Neighborhood segregation and black entrepreneurship

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HIGHLIGHTS

• We find that segregation has a positive effect on black entrepreneurship.
• We address neighborhood sorting by analyzing city averages.
• We address omitted variable bias by instrumenting segregation with railroad configurations.
• Our findings are important because entrepreneurship may decrease welfare and unemployment.
• Entrepreneurship is an important avenue out of poverty.

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ABSTRACT

We examine the causal effect of neighborhood segregation on black entrepreneurship. We address neighborhood sorting by analyzing city averages and omitted variable bias by instrumenting for segregation using historical railroad configurations. We find that segregation has a significant positive effect: a 10 percentage point increase in the dissimilarity index decreases the racial gap by about 3.3 percentage points. To minimize the effect of cross-city sorting, we use a narrower sample constructed from outcomes of young adults and find a similar effect. Our findings are important because historically, entrepreneurship has been an avenue out of poverty, and entrepreneurship has been promoted as a way to decrease welfare and unemployment.

1. Introduction

We consider the effect of segregation on black entrepreneurship—a relationship that has been studied previously but, to the best of our knowledge, not in a manner that renders a causal estimate.1 Understanding the racial gap in entrepreneurship is important because entrepreneurship is a source of wealth and employment as well as a critical channel of upward mobility (Fairlie and Robb, 2008; Quadrini, 1999). Racial disparities in business ownership could exacerbate wealth inequality along racial lines, and engender persistent intergenerational economic stagnation for minorities.

Neighborhood segregation does not necessarily lead to lower black entrepreneurship rates especially if there are positive spillovers that result from within-group mixing among income classes (Higgs, 1977) or from market segmentation whereby black entrepreneurs serve the needs of black customers that are not pursued by white-owned businesses (Brimmer, 1997). However, if racial segregation results in a lack of positive role models or a deficient provision of local public goods, then it is plausible that segregation could curb black entrepreneurship.2

1 See, for example, Massey and Denton (1993) and Bogan and William (2008). Cutler and Glaeser (1997), Card and Rothstein (2007), and Ananat (2011) examine the effect of segregation on other black outcomes such as employment status, test scores, and income.

2 Wilson (1996) observed that “Segregation in ghettos exacerbates employment problems because it leads to weak informal employment networks and contributes to social isolation of individuals and families, thereby reducing their chances of acquiring the human capital skills, that facilitate mobility in society. Since no other
Establishing the causal effect of segregation on entrepreneurship is complicated by two primary concerns. First, individuals may self-select into neighborhoods. For example, more enterprising blacks may choose to locate in less segregated neighborhoods. To mitigate this concern, following Cutler and Glaeser (1997) and Card and Rothstein (2007), we average outcomes to the city level for non-blacks and blacks and take the difference to eliminate any city-wide variables that affect the two groups equally. To account for possible city-level unobservables that affect the groups differently, we include city characteristics in the specification of the remaining error term. The second concern is that omitted variable bias could arise from unobservable city-level attributes that affect both segregation and mean economic outcomes. We instrument for segregation using the Railroad Division Index (RDI) of Ananat (2011). Ananat argues that the extent to which a city was subdivided by nineteenth-century railroad tracks, which subsequently served as natural enclave boundaries, influenced how segregated a city became when large inflows of blacks moved during the Great Migration.

Addressing these concerns, we find strong evidence that greater neighborhood segregation increases relative black entrepreneurship. A 10 percentage point increase in the dissimilarity index, an index that measures the level of neighborhood segregation, increases the rate of black entrepreneurship by 3.3 percentage points relative to the rate of non-blacks. To minimize the influence of cross-city sorting, we also estimate the segregation effect using a narrower sample constructed from the outcomes of young adults. This narrower sample mitigates the influence of sorting since young adults have a shorter window to change cities, and the likelihood of such moves is conceivably low. Using this sample we find that a 10 percentage point increase in the dissimilarity index increases the rate of black entrepreneurship by 2.8 percentage points relative to the rate of non-blacks.

2. Model

The outcome of individual $i$ of racial group $j$ living in city $c$ is determined by

$$ Y_{ijc} = X_{ijc} \alpha + R_{ijc} \delta + \epsilon_{ijc}, \quad (1) $$

where

$$ Y_{ijc} = \begin{cases} 1 & \text{if the individual is an entrepreneur} \\ 0 & \text{if the individual is employed by others} \end{cases} $$

$X_{ijc}$ is a vector of observed individual characteristics, and $R_{ijc}$ is the fraction of blacks in $i$'s neighborhood. $\delta$ is the parameter of interest. It measures the effect of neighborhood segregation on entrepreneurship. The error $\epsilon_{ijc}$ has two components. One component is common to individuals in racial group $j$ living in city $c$, $\epsilon_{jc}$, the other component is an individual-specific error with mean 0 for each racial group living in each city, $\xi_{ijc}$.

Following Cutler and Glaeser (1997) and Card and Rothstein (2007), we average outcomes of each racial group to the city level which removes $\xi_{ijc}$ from the model and eliminates the effect of non-random sorting of households into neighborhoods within a given city. Taking the average of (1), we have

$$ Y_{jc} = X_{jc} \alpha + R_{jc} \delta + \epsilon_{jc}, \quad (2) $$

Here, $Y_{jc}$ is the entrepreneurship rate of group $j$ in city $c$, $X_{jc}$ are the mean characteristics of racial group $j$ living in city $c$, and $R_{jc}$ is the average fraction of black neighbors in group $j$ living in city $c$.

We then take the difference between racial groups within a city to eliminate any city-wide variables that affect the two racial groups equally:

$$ \Delta Y_c = \Delta X_c \alpha + \Delta R_c \delta + \Delta \epsilon_c, \quad (3) $$

where $\Delta Y_c = Y_{2c} - Y_{1c}$, $\Delta X_c = X_{2c} - X_{1c}$, and $\Delta \epsilon_c = u_{2c} - u_{1c}$. $\Delta R_c$ is the dissimilarity index, a measure of the level of segregation in city $c$.

To account for any possible unobserved differences between non-blacks and blacks at the city level, we include city characteristics in the specification of $\Delta \epsilon_c$. That is,

$$ \Delta \epsilon_c = F_c \psi + \nu_c $$

where $F_c$ are city characteristics and $\nu_c$ contains the remaining unobserved differences between non-blacks and blacks in city $c$.

The model to be estimated is then

$$ \Delta Y_c = \Delta X_c \alpha + \Delta R_c \delta + F_c \psi + \nu_c. \quad (4) $$

As mentioned previously, omitted variable bias could still be present. We address this by instrumenting for neighborhood segregation following Ananat (2011). There is also the possibility of cross-city sorting. We address this in a robustness test by estimating the model on a sample created from outcomes of young adults, following the approach of Cutler and Glaeser (1997).

3. Data

Our data comes from four sources. Data on entrepreneurship and individual characteristics are from the 5-percent Public Use Microdata Sample Files (PUMS) of the 2000 Census. City characteristics for 2000 were downloaded using American FactFinder. The 2000 dissimilarity index was downloaded from the archived web page of Jacob Vigdor.3 Our instrument for the dissimilarity index, the Railroad Division Index (RDI), and 1910 and 1920 city characteristics are from Ananat (2011).4

3.1. 2000 Census Data

Using the 2000 Census data, for our primary analysis, we computed entrepreneurship rates and average characteristics by Metropolitan Statistical Area (MSA) and by racial group using US-citizen heads of household ages 18–65 who were not in school or the armed forces. A person is an entrepreneur if the PUMS class-of-worker variable indicates that the person worked for their own enterprise, and is not an entrepreneur if the person worked for someone else as an employee. Our dependent variable is the difference in mean entrepreneurship rates of non-blacks and blacks. For ease of exposition, we will henceforth refer to this dependent variable as the racial gap.

We also estimated the model using a sample based on 18–25 year olds. This is meant to minimize the effect of cross-city sorting since young adults have had only a short period in which to change residence. Moreover, we assign individuals to their MSA of residence five years before being interviewed to capture the effect of segregation when peer influences are presumably strongest. This subsample contains fewer observations since some MSAs do not contain observations from any young black individuals.

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3 http://trinity.aas.duke.edu/~jvigdor/segregation.
4 The data was downloaded from the AEA webpage (https://www.aeaweb.org/articles?id=10.1257/app.3.2.34).
3.2. Dissimilarity Index

To measure the level of segregation within a MSA, we use the standard dissimilarity index:

\[
\text{Dissimilarity Index} = \frac{1}{2} \sum_{i=1}^{N} \left( \frac{\text{black}_i}{\text{black}_{\text{total}}} - \frac{\text{nonblack}_i}{\text{nonblack}_{\text{total}}} \right).
\]

Here, \( i \) identifies a census tract within an MSA. The dissimilarity index measures the fraction of blacks that would have to move to a different census tract in order for the proportion black in each neighborhood to equal the proportion black in the metropolitan area as a whole. If blacks were evenly distributed throughout the city, the index would be zero. If blacks and non-blacks were fully segregated, i.e., each census tract consisted of either no blacks or all blacks, then the index would be one. (In the analysis below, the index is on a 0–100 scale.)

3.3. Railroad Division Index (RDI)

We instrument for the dissimilarity index using the Railroad Division Index (RDI) of Ananat (2011). The RDI measures the extent to which a metropolitan area is subdivided by railroad tracks laid in the 19th century. It is computed as

\[
\text{RDI} = 1 - \sum_{i} \left( \frac{\text{AREA}_{\text{neighborhood}}}{\text{AREA}_{\text{total}}} \right)^2.
\]

An undivided city would have a single contiguous neighborhood and an RDI value of 0. A city that was infinitely subdivided by railroad tracks so that each subdivision had an area of near 0 would have an RDI value of 1.

Our identification strategy assumes that RDI in the 19th century does not predict the gap in non-black and black entrepreneurship in 2000, except through its effect on segregation. The primary motivation for this instrument, as argued by Ananat (2011), is that railroad tracks define spatial boundaries and highly subdivided cities tended to become more significantly segregated during the Great Migration (around 1915–1950). Our first-stage regression results in Table A.2 in the Online Appendix do indeed show that RDI significantly predicts 2000 segregation levels. There are two primary reasons however why our identification strategy could fail. First, RDI could be correlated with historical city characteristics which may have impacted later entrepreneurship. \(^5\) Second, the instrument may have had a direct impact on the early demographics and economic characteristics of each city.

Table 1 contains the results of six regressions measuring the relationship between RDI and various city characteristics in 1910, one decade after the end of major railroad construction and before the Great Migration. It shows that RDI is not related to any of these characteristics, indicating that railroad configuration was not driven by local economic or social characteristics.\(^6\)

To address the possibility that subsequent to the first wave of the Great Migration, people may have sorted themselves non-randomly based on RDI, we examined six human capital characteristics of cities as well as population density (to account for early urbanization economies) in 1920. The results in Table A.1 in the Online Appendix show that RDI is not correlated with these early city population characteristics that might also affect later entrepreneurship rates.\(^7\)

3.4. Summary statistics and other covariates

Our primary sample consists of 98 MSA-level observations. On average, the entrepreneurship rate was 12\% for non-blacks compared to 4\% for blacks, resulting in a racial gap of around 8 percentage points. The average metropolitan area had a dissimilarity index value of 52 and an RDI value of 0.72.

In the full regression, we include average non-black/black differences in age and college degree attainment. Additionally, we include the following city characteristics: average track length per 100 square km,\(^8\) population, black population, Hispanic population, land area, natural log of median income, and manufacturing share of employment.

4. Results

Table 2 contains estimates of our model. The top panel contains estimates from the full sample. The bottom panel contains estimates from the narrower sample computed using individuals 18 to 25 years of age.\(^9\)

In both panels, columns (1) and (2) are OLS estimates of the model without covariates and with covariates, respectively. Columns (3) and (4) contain IV estimates without and with additional covariates.

The OLS and IV estimates in the top panel shows that segregation has a negative effect on the entrepreneurship gap between non-blacks and blacks, with and without additional controls. That is, more segregated metropolitan areas have a smaller racial gap in the entrepreneurship rate. Both set of IV estimates are significant at at least the 5\% level. The IV estimate of the effect of segregation in the full model reported in column (4) indicates that a 10 percentage point increase in the dissimilarity index decreases the racial gap by about 3.3 percentage points.

\(^5\) For example, the configuration of 19th century railroad tracks may be correlated with the connectedness of a city or the early industrial base and this could generate early selection and migration patterns that affect later entrepreneurship opportunities. We thank an anonymous referee for pointing this out.

\(^6\) These findings are similar to those in Ananat (2011). We re-estimated her regressions since our sample includes a smaller set of cities as entrepreneurship rates are not available for all the cities used in her analysis.

\(^7\) The Online Appendix also provides qualitative arguments, borrowed heavily from Ananat (2011), on why RDI is a valid instrument.

\(^8\) Track length is included as a regressor in all regressions to ensure that RDI is not capturing the amount of railroad track in an area.

\(^9\) Tables A.2, A.3, and A.4 of the Online Appendix contain the complete results of the first and second stage regressions.
Table 2
Entrepreneurship Regressions.

<table>
<thead>
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<th>Ages 18 to 65</th>
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<th>2SLS</th>
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<td>(2)</td>
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<td>−0.054***</td>
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<td>Observations</td>
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<td>85</td>
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<table>
<thead>
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<th>2SLS</th>
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<tbody>
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<td>(1)</td>
<td>(2)</td>
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<tr>
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<tr>
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</table>

Notes: *** p < 0.01, ** p < 0.05, * p < 0.1. The dependent variable is the non-black/black difference in the average rate of entrepreneurship by MSA. In columns (3) and (4), Dissimilarity Index is instrumented with the Railroad Division Index of Ananat (2011). Additional covariates include non-black/black differences in average age and education, and total population, black population, Hispanic population, log median income, manufacturing share of employment, land area, and track length. Robust standard errors are in parentheses.

In the lower panel, we see that the estimates of the effect of segregation are reasonably similar when we restrict the sample to young adults, suggesting that cross-city selection may not be a significant problem in the wider sample. In the full specification in column (4), the IV estimate indicates that a 10 percentage point increase in the dissimilarity index causes a 2.8 percentage point decrease in the racial gap.

5. Conclusion

American cities are characterized by residential segregation, which the literature argues could lead to a positive or a negative effect on black entrepreneurship. Racial enclaves may create protected markets for black-owned businesses serving black consumers that are kept out of markets due to discrimination (a positive effect) or they could reflect racism-motivated socio-economic isolation and deprivation (a negative effect). We find that, in fact, segregation decreases the racial gap in entrepreneurship. Blacks living in more segregated cities are more likely to be self-employed compared to others. Prima facie, our results seem to suggest that segregated neighborhoods are conducive to market segmentation in which black businesses serve predominantly black customers. Alternatively, our results could be a reflection of the effectiveness of place-based affirmative action programs as suggested by Boston (1999). Further research in this area is warranted.

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Appendix A. Supplementary material

Supplementary material related to this article can be found online at http://dx.doi.org/10.1016/j.econlet.2017.02.025.

References


10 See the discussion on the origins and evolution of black businesses in Brimmer (1997).

11 An example is the well-studied “Atlanta Plan” initiated in 1974 by Mayor Maynard Jackson. It paved the way for black entrepreneurs to break into city contracting through mandated set-asides funds to support minority-owned businesses. Total minority procurement increased from 0.13% in 1973 to 38.5% in 1978, the end of Jackson’s first term (Boston, 1999, p. 15). The gain in minority procurement arguably created a momentum: Atlanta’s black-owned businesses nearly doubled from 11,804 to 23,488 between 1987 and 1992. (Boston, 1999, p. 18).